

*CONVERGENCE RATES FOR WEIGHTED SUMS
OF RANDOM VARIABLES WITH RANDOM INDICES*

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1. Introduction and preliminaries. Let $\{a_{ik}\}$ and $\{b_{ik}\}$ for $i, k = 1, 2, \dots$ be double sequences of real numbers and let $\{X_k, k \geq 1\}$ be a sequence of not necessarily identically distributed random variables, defined on a probability space (Ω, \mathcal{A}, P) . Limit properties of sums

$$(1) \quad S_i = \sum_{k=1}^{\infty} a_{ik}(X_k - b_{ik}) \quad \text{as } i \rightarrow \infty$$

have been investigated in various papers, e.g. in [1]-[3] and [5]-[8]. It was proved that the sums of form (1) have many properties similar to those for random variables

$$(2) \quad T_i = \frac{1}{i} \sum_{k=1}^i X_k \quad \text{as } i \rightarrow \infty.$$

The aim of this paper is to give sufficient conditions for convergence in probability to zero of sums

$$(3) \quad S_{N_n} = \sum_k a_{N_n k}(X_k - b_{N_n k}) \quad \text{as } n \rightarrow \infty,$$

where $\{N_n, n \geq 1\}$ is a sequence of positive integer-valued random variables defined on the same probability space (Ω, \mathcal{A}, P) .

We assume that for every $n \geq 1$ and $k \geq 1$ the random variables N_n and X_k are independent. Moreover, we suppose that sequences $\{N_n, n \geq 1\}$ and $\{a_{ik}, i, k \geq 1\}$ are such that either

$$(4) \quad \mathbb{E} \sup_j |a_{N_n j}| \sum_k |a_{N_n k}|^t = o\left(\mathbb{E} \sum_k |a_{N_n k}|^t\right) \quad \text{as } n \rightarrow \infty$$

or

$$(5) \quad \mathbb{E} \left(\sum_k |a_{N_n k}|^t \right)^2 = o\left(\mathbb{E} \sum_k |a_{N_n k}|^t\right) \quad \text{as } n \rightarrow \infty,$$

where $0 < t < \infty$.

Further on, we suppose that $\{\varrho_i(t), i \geq 1\}$, $t \in R$, is a sequence of non-negative real numbers such that

$$(6) \quad \mathbb{E} \sum_k |a_{N_n k}|^t \leq \mathbb{E} \varrho_{N_n}(t), \quad n \geq 1.$$

To abbreviate the notation, set

$$F_k(y) = P[X_k < y], \quad F'_k(y) = P[X_k - \mathbb{E}X_k < y], \quad k \geq 1,$$

if $\mathbb{E}X_k$ exists, and put

$$F(y) = \sup_k P[|X_k| \geq y], \quad F'(y) = \sup_k P[|X_k - \mathbb{E}X_k| \geq y].$$

2. Results.

THEOREM 1. *Let $\{X_k, k \geq 1\}$ be a sequence of random variables and $\{N_n, n \geq 1\}$ a sequence of positive integer-valued random variables both defined on a probability space (Ω, \mathcal{A}, P) and such that, for every $k \geq 1$ and $n \geq 1$, X_k and N_n are independent.*

1° *If $0 < t < 1$, $y^t F(y) \leq M < \infty$ for all $y \geq 0$, and (6) holds, then*

$$(O) \quad P[|S_{N_n}| > \varepsilon] = O(\mathbb{E} \varrho_{N_n}(t)) \quad \text{as } n \rightarrow \infty,$$

with $b_{N_n k} = 0$ a.s.

2° *If $0 < t < 1$, $y^t F(y) \rightarrow 0$ as $y \rightarrow \infty$, and (4) and (6) hold, then*

$$(o) \quad P[|S_{N_n}| > \varepsilon] = o(\mathbb{E} \varrho_{N_n}(t)) \quad \text{as } n \rightarrow \infty,$$

with $b_{N_n k} = 0$ a.s.

In the following theorems we assume that $\{X_k, k \geq 1\}$ is a sequence of independent random variables, and a sequence $\{N_n, n \geq 1\}$ satisfies the assumptions of Theorem 1.

THEOREM 2. 1° *If $t = 1$, $yF(y) \leq M < \infty$ for all $y \geq 0$, inequality (6) is satisfied, and*

$$(7) \quad \overline{\lim}_{T \rightarrow \infty} \sup_k \left| \int_{-T}^T y dF_k(y) \right| < \infty,$$

then (O) holds with $b_{N_n k} = 0$ a.s.

2° *If $t = 1$, $yF(y) \rightarrow 0$ as $y \rightarrow \infty$, and (5), (6), and (7) are satisfied, then (o) holds with $b_{N_n k} = 0$ a.s.*

THEOREM 3. 1° *If $t = 1$, $yF(y) \leq M < \infty$, and (6) is satisfied, then (O) holds with*

$$(8) \quad b_{N_n k} = \begin{cases} |a_{N_n k}|^{-1} \int y dF_k(y) & \text{if } N_n \notin [i: a_{ik} = 0], \\ -|a_{N_n k}|^{-1} & \\ 0 & \text{if } N_n \in [i: a_{ik} = 0] \text{ a.s.} \end{cases}$$

2° If $t = 1$, $yF(y) \rightarrow 0$ as $y \rightarrow \infty$, and (4) and (6) are satisfied, then (o) holds with $b_{N_n k}$ given by (8).

THEOREM 4. 1° If $1 < t < 2$, $y^t F(y) \leq M < \infty$, and (6) is satisfied, then (O) holds with $b_{N_n k} = EX_k$ a.s.

2° If $1 < t < 2$, $y^t F(y) \rightarrow 0$ as $y \rightarrow \infty$, and (4) and (6) are satisfied, then (o) holds with $b_{N_n k} = EX_k$ a.s.

THEOREM 5. 1° If $t = 2$, $y^2 F(y) \leq M < \infty$, inequality (6) is satisfied and, moreover,

$$\exists 0 < \lambda < \infty \quad \forall 0 < \rho, \mu \leq \max[8, 4/\lambda] + 2,$$

$$(9) \quad E \left(\sum_k a_{N_n k}^2 \right)^\rho \left(\sum_j a_{N_n j}^2 |\log |a_{N_n j}|| \right)^\mu = O \left(E \left(\sum_k a_{N_n k}^2 \right)^{\rho + \lambda \mu} \right) \quad \text{as } n \rightarrow \infty,$$

then (O) holds with $b_{N_n k} = EX_k$ a.s.

2° If $t = 2$, $y^2 F(y) \rightarrow 0$ as $y \rightarrow \infty$, and (5), (6), and (9) are satisfied, then (o) holds with $b_{N_n k} = EX_k$ a.s.

THEOREM 6. 1° If $t > 2$, $y^t F(y) \leq M < \infty$, inequality (6) is satisfied and, moreover,

$$\exists 0 < a < \infty \quad \forall a < \beta < 6ta + 2 \quad \forall 0 < \gamma < 6ta + 2,$$

$$(10) \quad E \left(\sum_k a_{N_n k}^2 \right)^\beta \left(\sum_k |a_{N_n k}|^t \right)^\gamma = O \left(E \left(\sum_k |a_{N_n k}|^t \right)^{\beta/a + \gamma} \right)$$

as $n \rightarrow \infty$, then (O) holds with $b_{N_n k} = EX_k$ a.s.

2° If $t > 2$, $y^t F(y) \rightarrow 0$ as $y \rightarrow \infty$, and (5), (6), and (10) are satisfied, then (o) holds with $b_{N_n k} = EX_k$ a.s.

LEMMA 1. The following conditions are equivalent:

$$(A) \quad \begin{cases} \text{There exists a random variable } X \text{ such that} \\ (a_1) \quad \forall k \in N, P[|X_k| \geq y] \leq P[|X| \geq y], \\ (a_2) \quad E|X|^t \leq M_1 < \infty. \end{cases}$$

$$(B) \quad \begin{cases} (b_1) \quad F(y) \rightarrow 0 \text{ as } y \rightarrow \infty, \\ (b_2) \quad \int_0^\infty y^t |dF(y)| \leq M_2 < \infty. \end{cases}$$

THEOREM 7. *If $t = 1$, $F(y) \rightarrow 0$ as $y \rightarrow \infty$, (4) and (6) are satisfied and, moreover,*

$$(11) \quad \int_0^{\infty} y |dF(y)| \leq M < \infty,$$

then (o) holds with $b_{N_n k} = EX_k$ a.s.

By Lemma 1, if $\{X_k, k \geq 1\}$ is a sequence of independent random variables uniformly bounded by a random variable X such that $E|X| < \infty$, we have

THEOREM 7'. *If $t = 1$, and (4) and (6) are satisfied, then (o) holds with $b_{N_n k} = EX_k$ a.s.*

In all proofs we put $P[N_n = i] = p_{in}$. Summations in (1) and (3) may be taken only over those values of i for which $a_{ik} = 0$. Integrals will be Lebesgue-Stieltjes ones. By C we shall denote different, in general, positive absolute constants.

Proof of Theorem 1. By Theorem 1 of [8], for any given $\varepsilon > 0$ we have

$$P[|S_i| > \varepsilon] \leq C \sum_k |a_{ik}|^t,$$

and since X_k and N_n are independent for any $k \geq 1$ and $n \geq 1$, we see that 1° holds.

To prove 2° let us put

$$I = [i: \sup_k |a_{ik}| \leq \eta],$$

where $\eta > 0$ will be fixed later. Then for an arbitrary $\varepsilon > 0$ we have

$$P[|S_{N_n}| > \varepsilon] = \sum_{i \notin I} P[|S_i| > \varepsilon] p_{in} + \sum_{j \in I} P[|S_j| > \varepsilon] p_{jn}.$$

Now let τ be any given positive number. For sufficiently small fixed η , by Theorem 2 of [8] we obtain

$$\sum_{i \in I} P[|S_i| > \varepsilon] p_{in} \leq \frac{\tau}{2} E \varrho_{N_n}(t).$$

Further, by (4), for sufficiently large n we have

$$\sum_{j \notin I} P[|S_j| > \varepsilon] p_{jn} \leq \frac{1}{\eta^2} \sum_{j \notin I} \sup_i |a_{ji}| \sum_k |a_{jk}|^t p_{jn} \leq \frac{\tau}{2} E \varrho_{N_n}(t),$$

which proves 2°.

Using Theorems 1b and 2b of [3] and the considerations given in the proof of Theorem 1 one can easily establish the statements of Theorem 2.

Proof of Theorem 3. 1° Since random variables X_k are independent of N_n , we have the inequality

$$(12) \quad \begin{aligned} \mathbb{P}[|S_{N_n}| > \varepsilon] &\leq \sum_i \sum_k \mathbb{P}[|a_{ik} X_k| > 1] p_{in} + \\ &+ \sum_i \mathbb{P}\left[\left|\sum_k a_{ik}(Y_{ik} - b_{ik})\right| > \varepsilon\right] p_{in}, \end{aligned}$$

where $Y_{ik} = X_k \mathcal{I}[|a_{ik} X_k| \leq 1]$, and $\mathcal{I}[A]$ denotes the indicator of A .

The first term on the right-hand side of (12) can be bounded as follows:

$$(13) \quad \begin{aligned} \sum_i \sum_k \mathbb{P}[|a_{ik} X_k| > 1] p_{in} &\leq \sum_i \sum_k F\left(\frac{1}{|a_{ik}|}\right) p_{in} \\ &\leq M \sum_i \sum_k |a_{ik}|^t p_{in} = O(\mathbb{E} \varrho_{N_n}(t)). \end{aligned}$$

For the second term in (12), since $b_{ik} = \mathbb{E} Y_{ik}$, by Chebyshev's inequality we have the estimation

$$(14) \quad \begin{aligned} \sum_i \mathbb{P}\left[\left|\sum_k a_{ik}(Y_{ik} - b_{ik})\right| > \varepsilon\right] p_{in} \\ &\leq C \sum_i \sum_k a_{ik}^2 \mathbb{E} Y_{ik}^2 p_{in} \leq C \sum_i \sum_k \int_0^{|a_{ik}|^{-1}} 2yF(y) dy p_{in} \\ &\leq C \sum_i \sum_k |a_{ik}| p_{in} = O(\mathbb{E} \varrho_{N_n}(t)). \end{aligned}$$

To prove statement (o) we make the assumption

$$\forall \tau > 0 \exists \eta > 0 \forall y \geq \frac{1}{\eta}, \quad y^t F(y) \leq \frac{\tau}{2},$$

estimating the terms on the right-hand side of (12). It can be seen that, for

$$i \in I = [i: \sup_k |a_{ik}| \leq \eta],$$

we have

$$(15) \quad \sum_i \sum_k \mathbb{P}[|a_{ik} X_k| > 1] p_{in} \leq \sum_i \sum_k F\left(\frac{1}{|a_{ik}|}\right) p_{in} \leq \frac{\tau}{2} \mathbb{E} \varrho_{N_n}(t).$$

If $j \notin I$, then, using (4),

$$(16) \quad \begin{aligned} \sum_j \sum_k \mathbb{P}[|a_{jk} X_k| > 1] p_{jn} \\ &\leq C \sum_j \sup_k |a_{jk}| \sum_i |a_{ji}|^t \sup_{y>0} y^t F(y) p_{jn} \leq \frac{\tau}{2} \mathbb{E} \varrho_{N_n}(t). \end{aligned}$$

Now, let us observe that

$$\forall \tau > 0 \exists \eta > 0 \forall T \geq \frac{1}{\eta}, \frac{1}{T} \int_0^T 2yF(y) dy \leq \frac{\tau \varepsilon^2}{2}.$$

By this fact and considerations which have led to (14), we have, for $i \in I$,

$$(17) \quad \sum_i \mathbb{P} \left[\left| \sum_k a_{ik}(Y_{ik} - b_{ik}) \right| > \varepsilon \right] p_{in} \leq \frac{\tau}{2} \mathbb{E} \varrho_{N_n}(t).$$

If $j \notin I$, then by (4) we get

$$(18) \quad \sum_j \mathbb{P} \left[\left| \sum_k a_{jk}(Y_{jk} - b_{jk}) \right| > \varepsilon \right] p_{jn} \\ \leq C \sum_j \sup_i |a_{ji}| \sum_k |a_{jk}| p_{jn} \sup_{T>0} \frac{1}{T} \int_0^T 2yF(y) dy \leq \frac{\tau}{2} \mathbb{E} \varrho_{N_n}(t).$$

Inequalities (15)-(18) together complete the proof of 2°.

Proof of Theorem 4. By Theorem 1c of [3] we know that

$$\mathbb{P}[|S_i| > \varepsilon] \leq C \sum_k |a_{ik}|^t$$

and we see that the first assertion of Theorem 4 can be obtained in the same way as 1° of Theorem 1.

To prove 2° we need the inequality

$$(19) \quad \sum_i \mathbb{P}[|S_i| > 2\varepsilon] p_{in} \leq \sum_i \sum_k \mathbb{P}[|a_{ik}(X_k - \mathbb{E}X_k)| > 1] p_{in} \\ \leq \sum_i \mathbb{P} \left[\left| \sum_k a_{ik} \mathbb{E}Z_{ik} \right| > \varepsilon \right] p_{in} + \sum_i \mathbb{P} \left[\left| \sum_k a_{ik}(Z_{ik} - \mathbb{E}Z_{ik}) \right| > \varepsilon \right] p_{in},$$

where

$$Z_{ik} = (X_k - \mathbb{E}X_k) \mathcal{I} [|a_{ik}(X_k - \mathbb{E}X_k)| \leq 1].$$

Taking into account (15) and (16) and putting F' instead of F we see that

$$\sum_i \mathbb{P}[|a_{ik}(X_k - \mathbb{E}X_k)| > 1] p_{in} = o(\mathbb{E} \varrho_{N_n}(t)).$$

Now we observe that

$$(20) \quad |\mathbb{E}Z_{ik}| \leq \int_{|a_{ik}|^{-1}}^{\infty} y |dF'(y)| = \frac{1}{|a_{ik}|} F' \left(\frac{1}{|a_{ik}|} \right) + \int_{|a_{ik}|^{-1}}^{\infty} F'(y) dy.$$

Using (20), we have

$$(21) \quad \sum_i P \left[\left| \sum_k a_{ik} \mathbf{E} Z_{ik} \right| > \varepsilon \right] p_{in} \\ \leq C \sum_i \sum_k F' \left(\frac{1}{|a_{ik}|} \right) p_{in} + C \sum_i \sum_k |a_{ik}| \int_{|a_{ik}|^{-1}}^{\infty} F'(y) dy.$$

By the same arguments as for (15) and (16), we get

$$\sum_i \sum_k F' \left(\frac{1}{|a_{ik}|} \right) p_{in} = o(\mathbf{E} \varrho_{N_n}(t)).$$

Since, by the assumptions of Theorem 4,

$$\forall \tau > 0 \exists \eta > 0 \forall |a_{ik}| \leq \eta, \quad \int_{|a_{ik}|^{-1}}^{\infty} F'(y) dy \leq \frac{\tau}{2} \varepsilon |a_{ik}|^{t-1},$$

for $i \in I$ we have

$$(22) \quad \sum_i \sum_k |a_{ik}| \int_{|a_{ik}|^{-1}}^{\infty} F'(y) dy p_{in} \leq \frac{\tau}{2} \mathbf{E} \varrho_{N_n}(t).$$

If $j \notin I$, then by (4)

$$(23) \quad \sum_j \sum_k |a_{jk}| \int_{|a_{jk}|^{-1}}^{\infty} F'(y) dy p_{jn} \leq C \sum_j \sum_k |a_{jk}| \int_{|a_{jk}|^{-1}}^{\infty} y^{-t} dy p_{jn} \\ \leq C \sum_j \sum_k |a_{jk}|^t \sup_l |a_{jl}| p_{jn} \leq \frac{\tau}{2} \mathbf{E} \varrho_{N_n}(t).$$

To estimate the last term of (19), let us first observe that under the assumptions of Theorem 4

$$\forall \tau > 0 \exists \eta > 0 \forall y \geq \frac{1}{\eta}, \quad 2yF'(y) \leq \tau y^{1-t} (2-t).$$

Now set

$$I = [i: \sup_k |a_{ik}| \leq \min[\eta, \tau^{1/(2-t)}]].$$

If $i \in I$, then

$$(24) \quad \sum_i P \left[\left| \sum_k a_{ik} (Z_{ik} - \mathbf{E} Z_{ik}) \right| > \varepsilon \right] p_{in} \leq \frac{1}{\varepsilon^2} \sum_i \sum_k a_{ik}^2 \int_0^{|a_{ik}|^{-1}} 2yF'(y) dy p_{in} \\ \leq C \sum_i \sum_k a_{ik}^2 \int_0^{1/\eta} 2y dy p_{in} + \tau C \sum_i \sum_k a_{ik}^2 |a_{ik}|^{t-2} p_{in} \\ \leq C \sum_i \sum_k |a_{ik}|^t \sup_l |a_{il}|^{2-t} p_{in} + \tau C \sum_i \sum_k |a_{ik}|^t p_{in} \leq \tau C \mathbf{E} \varrho_{N_n}(t).$$

If $j \notin I$, then, by (4),

$$(25) \quad \sum_j \mathbb{P} \left[\left| \sum_k a_{jk} (Z_{jk} - \mathbb{E}Z_{jk}) \right| > \varepsilon \right] p_{jn} \leq C \sum_j \sum_k a_{jk}^2 \int_0^{|a_{jk}|^{-1}} y^{1-t} dy p_{jn} \\ \leq C \sum_j \sum_k |a_{jk}|^t \sup |a_{jl}| p_{jn} \leq \tau \mathbb{E} \varrho_{N_n}(t).$$

Thus we have proved Theorem 4.

Proof of Theorem 5. First we observe that, by the inequality given in [1],

$$(26) \quad \mathbb{P}[|\mathcal{S}_{N_n}| > 3\varepsilon] \leq \sum_i \sum_k \mathbb{P}[|a_{ik}(X_k - \mathbb{E}X_k)| > \varepsilon] p_{in} + \\ + \sum_i \sum_{j \neq k} \mathbb{P}[|a_{ij}(X_j - \mathbb{E}X_j)| > \delta_i] \mathbb{P}[|a_{ik}(X_k - \mathbb{E}X_k)| > \delta_i] p_{in} + \\ + \sum_i \mathbb{P} \left[\left| \sum_k a_{ik} \mathbb{E}Z_{ik} \right| > \varepsilon \right] p_{in} + \sum_i \mathbb{P} \left[\left| \sum_k a_{ik} (Z_{ik} - \mathbb{E}Z_{ik}) \right| > \varepsilon \right] p_{in},$$

where

$$\delta_i = \left(\sum_k a_{ik}^2 \right)^{1/4} \quad \text{and} \quad Z_{ik} = (X_k - \mathbb{E}X_k) \mathcal{I}[|a_{ik}(X_k - \mathbb{E}X_k)| \leq \delta_i].$$

Under the assumption $y^2 F(y) \leq M < \infty$ (see [3]) we have

$$\sum_i R_i p_{in} = \sum_i \sum_k \mathbb{P}[|a_{ik}(X_k - \mathbb{E}X_k)| > \varepsilon] p_{in} + \\ + \sum_i \sum_{j \neq k} \mathbb{P}[|a_{ij}(X_j - \mathbb{E}X_j)| > \delta_i] \mathbb{P}[|a_{ik}(X_k - \mathbb{E}X_k)| > \delta_i] p_{in} + \\ + \sum_i \mathbb{P} \left[\left| \sum_k a_{ik} \mathbb{E}Z_{ik} \right| > \varepsilon \right] p_{in} \leq C \sum_i \sum_k a_{ik}^2 p_{in}$$

and, consequently,

$$(27) \quad \sum_i R_i p_{in} \leq C \sum_i \sum_k a_{ik}^2 p_{in} = O(\mathbb{E} \varrho_{N_n}(t)).$$

Thus we only need to bound the last term in (26). We follow the method used in [1]-[3].

Let us choose a positive integer ν such that

$$\max[8, 4/\lambda] < 2\nu \leq \max[8, 4/\lambda] + 2.$$

Using the estimations given in the proof of Theorem 1d of [3], we obtain

$$\begin{aligned}
 (28) \quad \sum_i P \left[\left| \sum_k a_{ik} (Z_{ik} - \mathbb{E}Z_{ik}) \right| > \varepsilon \right] p_{in} \\
 \leq \sum_i C \sum^* \sum^{**} \prod_{k=1}^a |a_{i\beta_k}|^{m_k} \mathbb{E} |Z_{i\beta_k} - \mathbb{E}Z_{i\beta_k}|^{m_k} p_{in} \\
 \leq C \sum_i \sum^* \delta_i^{2\nu-2a} \left(\sum_k a_{ik}^2 \mathbb{E}Z_{ik}^2 \right)^a p_{in},
 \end{aligned}$$

where the sum \sum^* is taken over all integers a, m_1, m_2, \dots, m_a such that $2 \leq m_k, k = 1, 2, \dots, a$, and $m_1 + m_2 + \dots + m_a = 2\nu$, and in the sum \sum^{**} subscripts $\beta_1, \beta_2, \dots, \beta_a$ run over the positive integers.

It is enough to consider the case $\delta_i > e^{-2M'}$, where $y^2 F'(y) \leq M' < \infty$ for all $y \geq 0$. Thus

$$\sum_k a_{ik}^2 \mathbb{E}Z_{ik}^2 \leq \sum_k a_{ik}^2 \left(2M' \log \frac{\delta_i}{|a_{ik}|} + 1 \right) \leq 2M' \sum_k a_{ik}^2 |\log |a_{ik}||.$$

Now by (9) and (28) we have

$$\begin{aligned}
 (29) \quad \sum_i P \left[\left| \sum_k a_{ik} (Z_{ik} - \mathbb{E}Z_{ik}) \right| > \varepsilon \right] p_{in} \\
 \leq C \sum_i \sum^* \left(\sum_k a_{ik}^2 \right)^{(\nu-a)/2} \left(\sum_k a_{ik}^2 |\log |a_{ik}|| \right)^a p_{in} \\
 \leq C \sum_i \sum^* \left(\sum_k a_{ik}^2 \right)^{(\nu-a)/a + \lambda a} p_{in}.
 \end{aligned}$$

The sum \sum^* is finite, since it depends only on ν .

Now, we prove that for an arbitrary real number $\lambda > 0$ and for $a, 0 < a \leq \nu$,

$$(30) \quad \frac{\nu - a}{2} + \lambda a > 2.$$

If $\lambda \geq \frac{1}{2}$, then

$$\frac{\nu + a(2\lambda - 1)}{2} > \frac{\nu}{2} > 2,$$

since $2\nu > \max[8, 4/\lambda]$. If $0 < \lambda < \frac{1}{2}$, then

$$\frac{\nu + a(2\lambda - 1)}{2} \geq \frac{\nu + \nu(2\lambda - 1)}{2} = \nu\lambda > 2.$$

Taking into account (30), we see that the last expression in (29) is $O(\mathbb{E}e_{N_n}(t))$, which completes the proof of 1°.

To prove 2° let us put

$$I = \left[i: \sum_k a_{ik}^2 < \eta \right], \quad \text{where } \eta > 0.$$

The consideration similar to that in the proof of Theorem 1, 2°, after using Theorem 2d of [3] and inequality (26), allows us to write

$$\sum_i R_i p_{in} = o(\mathbb{E} \varrho_{N_n}(t)).$$

Finally, by (28) - (30), we have proved 2°.

Proof of Theorem 6. 1° The first term on the right-hand side of (26) can be bounded as (13), changing F into F' .

The second term is less than

$$(31) \quad \sum_i \left(\sum_k |a_{ik}|^t \delta_i^{-t} \right)^2 \left\{ \left(\frac{\delta_i}{|a_{ik}|} \right)^t F' \left(\frac{\delta_i}{|a_{ik}|} \right) \right\}^2 p_{in}.$$

Now let us write

$$\delta_i = \max \left[\left(\sum_k |a_{ik}|^t \right)^{1/3t}, \left(\sum_k a_{ik}^2 \right)^{1/3t} \right].$$

It suffices to consider the case where

$$\sum_k |a_{ik}|^t < 1.$$

Then

$$\delta_i^{-2t} \leq \left(\sum_k |a_{ik}|^t \right)^{-2t/3t} \leq \left(\sum_k |a_{ik}|^t \right)^{-1}.$$

Now (31) is of the form

$$\sum_i \sum_k |a_{ik}|^t \{ \sup y^t F'(y) \}^2 = O(\mathbb{E} \varrho_{N_n}(t)),$$

sup being taken for

$$y \geq \left(\sum_k |a_{ik}|^t \right)^{-1/t}.$$

To bound the third term on the right-hand side of (26) we use the inequality

$$(32) \quad \left| \sum_k a_{ik} \mathbb{E} Z_{ik} \right| \leq C \delta_i^{1-t} \sum_k |a_{ik}|^t \leq C \left(\sum_k |a_{ik}|^t \right)^{(1+2t)/3t}$$

(for the proof of this property, see [3]). If

$$C \left(\sum_k |a_{ik}|^t \right)^{(1+2t)/3t} < \varepsilon,$$

then

$$\mathbb{P} \left[\left| \sum_k a_{ik} \mathbb{E} Z_{ik} \right| > \varepsilon \right] = 0.$$

On the other hand, we have

$$(33) \quad \sum_i \mathcal{P} \left[\left| \sum_k a_{ik} \mathbb{E} Z_{ik} \right| > \varepsilon \right] p_{in} \leq C \sum_i \sum_k |a_{ik}|^t p_{in} = O(\mathbb{E} \varrho_{N_n}(t)).$$

Now we are going to estimate the last term of (26). Let us fix an integer ν such that $6ta < 2\nu \leq 6ta + 2$, and a real number μ , $0 < \mu < t - 2$. By Markov's inequality and by (2.29), (2.32), and (2.33) of [2], we get

$$(34) \quad \begin{aligned} & \mathbb{P} \left[\left| \sum_k a_{ik} (Z_{ik} - \mathbb{E} Z_{ik}) \right| > \varepsilon \right] \\ & \leq C \sum^* \sum^{**} \prod_{k=1}^{a+b} |a_{i\beta_k}|^{m_k} \mathbb{E} |Z_{i\beta_k} - \mathbb{E} Z_{i\beta_k}|^{m_k} \\ & \leq C \sum^* \left(\sum_k a_{ik}^2 \right)^{\frac{m_1 + \dots + m_a}{2} + \frac{b\mu}{t-2}} \left(\sum_k |a_{ik}|^t \right)^{\frac{b(t-\mu-2)}{t-2}} \delta_i^{m_{a+1} + \dots + m_{a+b} - bt + b\mu}, \end{aligned}$$

where the sum \sum^* is taken over all positive integers a, b and $m_k, k = 1, 2, \dots, a + b$, such that

$$2 \leq m_k < t \quad \text{for } k = 1, 2, \dots, a,$$

and

$$t \leq m_k \quad \text{for } k = a + 1, a + 2, \dots, a + b \text{ and } m_1 + m_2 + \dots + m_{a+b} = 2\nu,$$

and the sum \sum^{**} is taken over all sets of positive integers $(\beta_1, \beta_2, \dots, \beta_{a+b})$.

As previously, it suffices to consider the case where

$$\sum_k |a_{ik}|^t < 1.$$

Then

$$\delta_i = \left(\sum_k a_{ik}^2 \right)^{1/3t}.$$

It is easy to verify that

$$(35) \quad \left(\sum_k a_{ik}^2 \right)^{\frac{m_1 + \dots + m_a}{2}} \delta_i^{-(m_1 + \dots + m_a)} \leq \left(\sum_k a_{ik}^2 \right)^{(m_1 + \dots + m_a) \left(\frac{1}{2} - \frac{1}{3t} \right)}$$

and

$$(36) \quad \left(\sum_k a_{ik}^2 \right)^{\frac{\mu}{t-2}} \left(\sum_k |a_{ik}|^t \right)^{\frac{t-\mu-2}{t-2}} \delta_i^{-t} \leq \left(\sum_k a_{ik}^2 \right)^{\frac{2}{3} \frac{\mu}{t-2}} \left(\sum_k |a_{ik}|^t \right)^{\frac{2}{3} \frac{t-\mu-2}{t-2}}.$$

Thus (34) takes the form

$$(37) \quad \begin{aligned} & \mathbb{P} \left[\left| \sum_k a_{ik} (Z_{ik} - \mathbb{E} Z_{ik}) \right| > \varepsilon \right] \\ & \leq C \sum^* \left(\sum_k a_{ik}^2 \right)^{(m_1 + \dots + m_a) \left(\frac{1}{2} - \frac{1}{3t} \right) + \frac{2b\mu}{3(t-2)} + \frac{2\nu + b\mu}{3t}} \left(\sum_k |a_{ik}|^t \right)^{\frac{5}{3} \frac{b(t-\mu-2)}{t-2}}. \end{aligned}$$

By (10) and (37), we get

$$(38) \quad \sum_i \mathbb{P} \left[\left| \sum_k a_{ik} (Z_{ik} - \mathbb{E}Z_{ik}) \right| > \varepsilon \right] p_{in} \\ \leq C \sum_i \sum_k^* \left(\sum_k |a_{ik}|^t \right)^{\frac{m_1 + \dots + m_a}{a} \left(\frac{1}{2} - \frac{1}{3t} \right) + \frac{2}{3a} \frac{b\mu}{t-2} + \frac{2\nu}{3ta} + \frac{b\mu}{3ta} + \frac{5}{3} \frac{b(t-\mu-2)}{t-2}} p_{in}.$$

Since the sum \sum^* is finite, it is enough to show that the exponent in (38) is greater than 1. This is true in view of the assumption $6t\alpha < 2\nu \leq 6t\alpha + 2$. Thus we have proved 1°.

To prove 2° we estimate the right-hand side of (26) under assumption (5).

Let us set

$$I = \left[i : \left(\sum_k |a_{ik}|^t \right) \leq \eta \right].$$

Putting F' instead of F into (15) and (16), we obtain statement (o) for the first term on the right-hand side of (26). By (31), we see that under the assumptions of Theorem 6, 2°, the second term of (26) is $o(\mathbb{E}e_{N_n}(t))$. For the proof we consider the set

$$I = \left[i : \left(\sum_k |a_{ik}|^t \right)^{1/t} \leq \eta \right],$$

where $y^t F'(y)$ is sufficiently small for $y \geq 1/\eta$.

To estimate the third term, we observe that

$$\mathbb{P} \left[\left| \sum_k a_{ik} \mathbb{E}Z_{ik} \right| > \varepsilon \right] = 0 \quad \text{if } C \left(\sum_k |a_{ik}|^t \right)^{(1+2t)/3t} < \varepsilon$$

and, in the opposite case,

$$\sum_i \mathbb{P} \left[\left| \sum_k a_{ik} \mathbb{E}Z_{ik} \right| > \varepsilon \right] p_{in} \leq C \sum_i \left(\sum_k |a_{ik}|^t \right)^2 p_{in} = o(\mathbb{E}e_{N_n}(t)).$$

The last term in (26) is $o(\mathbb{E}e_{N_n}(t))$, for the exponent in (38) is greater than 2.

Thus the proof is completed.

Proof of Lemma 1. (A) \Rightarrow (B). By (a₂) we have

$$M \geq \mathbb{E}|X|^t \geq \int_{-y}^y |x|^t d\mathbb{P}[X < x] + y^t \mathbb{P}[|X| \geq y] \\ \geq \int_{-y}^y |x|^t d\mathbb{P}[X < x] \rightarrow \mathbb{E}|X|^t \quad \text{as } y \rightarrow \infty.$$

Hence, by (a₁),

$$0 \leq y^t F(y) \leq y^t \mathbb{P}[|X| \geq y] \rightarrow 0 \quad \text{as } y \rightarrow \infty,$$

which proves (b₁). Assertion (b₂) follows from

$$\begin{aligned} M &\geq \int_{-\infty}^0 |y|^t dP[X < y] + \int_0^{\infty} y^t dP[X < y] \\ &= \int_0^{\infty} (1 - P[X < y] + P[X < -y]) dy^t \geq \int_0^{\infty} F(y) dy^t = \int_0^{\infty} y^t |dF(y)|. \end{aligned}$$

(B) ⇒ (A). Let

$$G(y) = \begin{cases} 1 & \text{if } y = 0, \\ F(N) & \text{if } N < y \leq N + 1. \end{cases}$$

We observe that 1 - G(y) is the distribution function of a random variable X. Then, for every k ∈ N,

$$P[|X_k| \geq y] \leq F(y) \leq 1 - G(y) = P[X \geq y]$$

and

$$\begin{aligned} E|X|^t &= \int_0^{\infty} y^t |dG(y)| = \sum_{k=1}^{\infty} k^t [F(k-1) - F(k)] \\ &\leq 1 + 2^t \sum_{k=2}^{\infty} (k-1)^t [F(k-1) - F(k)] \leq 1 + 2^t M < \infty. \end{aligned}$$

Thus we have proved Lemma 1.

Proof of Theorem 7. Let us estimate the right-hand side of (19).

The first term can be bounded as (15) and (16) putting only F' instead of F.

To estimate the second term let us observe that by (20) and Markov's inequality we have

$$(39) \quad \sum_i P \left[\left| \sum_k a_{ik} EZ_{ik} \right| > \varepsilon \right] p_{in} \leq C \sum_i \sum_k |a_{ik}| \int_{1/|a_{ik}|}^{\infty} y |dF'(y)| p_{in}.$$

By the assumption of Theorem 7 we see that

$$\int_T^{\infty} y |dF'(y)| \rightarrow 0 \quad \text{as } T \rightarrow \infty.$$

Therefore, using the method of the previous proofs, we see that (39) is o(Eϱ_{N_n}(t)).

The last term of (19) can be estimated similarly as (24) and (25). This completes the proof.

Theorem 7' follows from Theorem 7 and Lemma 1.

3. Concluding remarks.

1. One can observe that our theorems yield, among others, results contained in [1]-[3] and [5]-[8]. To see that it is enough to put N_n = n a. s.

In this particular case, Theorem 4 is stronger than the corresponding result of [3]. Theorem 3 did not appear earlier.

2. Let a_{ik} be such that (4) holds and

$$\sum_k |a_{ik}| \leq C.$$

For such real numbers, by Theorem 7', $S_{N_n} \rightarrow 0$ in probability as $n \rightarrow \infty$. This fact extends the result of Rohatgi [7].

3. Let

$$(40) \quad a_{ik} = \begin{cases} 1/i^r & \text{for } 1 \leq k \leq i, \\ 0 & \text{for } k > i, \end{cases}$$

where $r > 1/t$. In this case,

$$\mathbf{E} \sum_k |a_{N_n k}|^t = \sum_i i^{1-rt} p_{in} = \mathbf{E} N_n^{1-rt},$$

and under the assumptions of Theorem 1, 1°, and Theorem 2, 1°, we have

$$\mathbf{P} \left[\left| \sum_{k=1}^{N_n} X_k \right| > \varepsilon N_n^r \right] = O(\mathbf{E} N_n^{1-rt}).$$

Theorem 3, 1°, gives

$$\mathbf{P} \left[\left| \sum_{k=1}^{N_n} \left(X_k - \int_{-N_n}^{N_n} x dF_k(x) \right) \right| > \varepsilon N_n^r \right] = O(\mathbf{E} N_n^{1-r})$$

and by Theorem 4, 1°,

$$\mathbf{P} \left[\left| \sum_{k=1}^{N_n} (X_k - \mathbf{E} X_k) \right| > \varepsilon N_n^r \right] = O(\mathbf{E} N_n^{1-rt}).$$

To obtain similar results for $t \geq 2$ we give the following

LEMMA 2. If $a_i > 0$, $b_i > 0$, $p_i \geq 0$ for $i = 1, 2, \dots$, and

$$\sum_{i=1}^{\infty} a_i p_i < \infty, \quad \sum_{i=1}^{\infty} b_i p_i < \infty, \quad \text{and} \quad \sum_{i=1}^{\infty} p_i = 1,$$

then

$$\frac{\sum_{i=1}^{\infty} a_i p_i}{\sum_{i=1}^{\infty} b_i p_i} \leq \sup_i \frac{a_i}{b_i}.$$

From Lemma 2 we conclude that assumption (9) is satisfied for a_{ik} given by (40) and for $0 < \lambda < 1$. Similarly, if

$$r > \frac{1}{2} \quad \text{and} \quad a > \frac{rt-1}{2r-1},$$

then (10) is valid. Thus, if $t \geq 2$, then

$$P \left[\left| \sum_{k=1}^{N_n} (X_k - EX_k) \right| > \varepsilon N_n^r \right] = O(EN_n^{1-rt}), \quad r > \frac{1}{2}.$$

To characterize convergence in probability to zero of sums with random indices we need

LEMMA 3. *If $N_n \rightarrow \infty$ in probability as $n \rightarrow \infty$, then, for an arbitrary $s < 0$, $EN_n^s \rightarrow 0$ as $n \rightarrow \infty$.*

By Lemma 3, (4) is satisfied for

$$a_{ik} = \begin{cases} \left(\frac{1}{i}\right)^{1/t} & \text{for } 1 \leq k \leq i, \\ 0 & \text{for } k > i. \end{cases}$$

Hence, under the assumptions of Theorems 1, 3, and 4, if $N_n \rightarrow \infty$ in probability, then, for $0 < t < 1$, $t = 1$, and $1 < t < 2$, we get, respectively,

$$P \left[\left| \sum_{k=1}^{N_n} X_k \right| > \varepsilon N_n^{1/t} \right] \rightarrow 0 \quad \text{as } n \rightarrow \infty,$$

$$P \left[\left| \sum_{k=1}^{N_n} \left(X_k - \int_{-N_n}^{N_n} x dF_k(x) \right) \right| > \varepsilon N_n \right] \rightarrow 0 \quad \text{as } n \rightarrow \infty,$$

and

$$P \left[\left| \sum_{k=1}^{N_n} (X_k - EX_k) \right| > \varepsilon N_n^{1/t} \right] \rightarrow 0 \quad \text{as } n \rightarrow \infty.$$

4. It is known that if $Y_n \rightarrow 0$ in probability as $n \rightarrow \infty$, N_n, Y_1, Y_2, \dots are independent for every $n \geq 1$ and, moreover, $N_n \rightarrow \infty$ in probability, then $Y_{N_n} \rightarrow 0$ in probability as $n \rightarrow \infty$.

Theorems obtained in this paper give the conditions under which $Y_{N_n} \rightarrow 0$ in probability as $n \rightarrow \infty$ without the assumption that $Y_n \rightarrow 0$ in probability as $n \rightarrow \infty$. Furthermore, they furnish information about rates of convergence in probability.

Example. Let $\{X_k, k \geq 1\}$ be a sequence of independent identically distributed random variables such that $E|X_1| < \infty$ and $EX_1 = 0$. Suppose that, for $p = 1, 2, \dots$,

$$a_{ik} = \begin{cases} 1/i & \text{for } 1 \leq k \leq i, i = 2p, \\ 0 & \text{for } k > i, i = 2p, \\ 1 & \text{for } k = 1, i = 2p-1, \\ 0 & \text{for } k > 1, i = 2p-1. \end{cases}$$

One can observe that

$$\sup_k |a_{ik}| \rightarrow 0 \quad \text{as } i \rightarrow \infty$$

and we cannot use Theorem 1 of [7] to assert that

$$S_i = \sum_k a_{ik} X_k \rightarrow 0 \text{ in probability as } i \rightarrow \infty.$$

Thus we do not know if $S_i \rightarrow 0$ in probability as $i \rightarrow \infty$.

Now let $\{N_n, n \geq 1\}$ be a sequence of positive integer-valued random variables such that, for $p = 1, 2, \dots$,

$$P[N_n = 1] = \frac{1}{n},$$

$$P[N_n = i] = \begin{cases} \frac{n^{p-1} - (1/n)^{p-1}}{(p-1)!} \frac{1-1/n}{e^n - 1} & \text{for } i = 2p, \\ \frac{(1/n)^{p-1}}{(p-1)!} \frac{1-1/n}{e^n - 1} & \text{for } i = 2p+1. \end{cases}$$

We have

$$E \sum_k |a_{N_n k}| = 1 \quad \text{and} \quad E \sup_j |a_{N_n j}| \sum_k |a_{N_n k}| \rightarrow 0 \quad \text{as } n \rightarrow \infty.$$

Therefore, by Theorem 7', $S_{N_n} \rightarrow 0$ in probability as $n \rightarrow \infty$.

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