

ON MIESHALKIN-ROGOZIN THEOREM AND SOME
PROPERTIES OF THE SECOND KIND BETA
DISTRIBUTION*

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Abstract

The decomposition of the r.v. X with the beta second kind distribution in the form of finite (formula (9), Theorem 1) and infinity products (formula (17), Theorem 2 and form (21), Theorem 3) are presented. Next applying Mieshalkin – Rogozin theorem we receive the estimation of the difference of two c.d.f. $F(x)$ and $G(x)$ when $\sup |f(t) - g(t)|$ is known, improving the result of Gnedenko – Kolmogorov (formulae (23) and (24)).

Keywords: Mieshalkin – Rogozin theorem, result of Kolmogorov, Knar formula.

1999 Mathematics Subject Classification: 60E07, 62E15.

1. INTRODUCTION

We consider a random variable X with the second kind beta distribution

$$(1) \quad \beta_2(x|p, q) = \frac{x^{p-1}}{B(p, q) (x+1)^{p+q}}, \quad 0 \leq x < \infty, \quad p, q > 0.$$

The random variable $X = Z/(1 - Z)$ has this distribution when Z has the first kind beta distribution

$$g(z) = B^{-1}(p, q) z^{p-1} (1 - z)^{q-1}, \quad 0 \leq z \leq 1.$$

*41 XX. Mieshalkin – Rogozin theorem and estimation of the difference of two c.d.f. $|F(x) - G(x)|$

A quotient $X = Y_1/Y_2$ of two independent gamma random variables with a probability density function $f_{Y_j}(y_j|b_j) = \Gamma^{-1}(b_j) y_j^{b_j-1} \exp(-y_j)$, $b_j > 0$, $0 \leq y_j < \infty$, $j = 1, 2$, has also the same distribution with $b_1 = p$, $b_2 = q$. In [16] Podolski has defined distributions of products and quotients of the powers of independent random variables with the second kind beta distribution using Meijer's G -function [3,9.3] and their generalization; the functions $H_{p,q}^{m,n}$ of Saksena and Mathai [18].

To represent a random variable X with the density function (1) as the finite product we shall use the Mellin transform $M_X(s)$ in the form of Zolotarev [20]

$$M_X(s) = EX^s,$$

where s is complex number, which gives

$$(2) \quad M_X(s) = \Gamma(p+s)\Gamma(q-s)\Gamma^{-1}(p)\Gamma^{-1}(q), \quad -p < \operatorname{Re} s < q.$$

We can present the gamma functions in (2) as a finite product [3,8.335].

2. THE CASE OF FINITE PRODUCT

Applying the Knar formula [3,8.335]

$$(3) \quad \Gamma(nx) = (2\pi)^{\frac{1-n}{2}} n^{nx-0,5} \prod_{k=1}^n \Gamma\left(x + \frac{k-1}{n}\right), \quad \operatorname{Re} x > 0, \quad n \geq 2,$$

to the gamma function in (2), we rewrite (2) for $n = n_1$ and $k = k_1$, after some simplifications, as

$$(4) \quad \begin{aligned} M_X(s) &= \prod_{k_1=1}^{n_1} \Gamma((p+k_1-1+s)/n_1) \Gamma((q+k_1-1-s)/n_1) \cdot \\ &\cdot \Gamma^{-1}((p+k_1-1)/n_1) \cdot \Gamma^{-1}((q+k_1-1)/n_1) = \\ &= \prod_{k_1=1}^{n_1} g_{k_1}(s), \quad -p < \operatorname{Re} s < q, \quad n_1 \geq 2. \end{aligned}$$

From the formula (2) it follows that the Mellin transform of random variable X_{k_1} with the second kind beta distribution with parameters $(p + k_1 - 1)/n_1$ and $(q + k_1 - 1)/n_1$ differs from the factor in the right hand side of the formula (4) only by a numerical coefficient at s . Because of $M_{X^{1/n}} = E \left(X^{1/n} \right)^s$, it follows that each factor $g_{k_1}(s)$ of the product (4) is the Mellin transform of the random variable $X_{k_1}^{1/n_1}$, where the random variables X_{k_1} have as densities the functions $\beta_2(x|p_{k_1}, q_{k_1})$, where

$$(5) \quad p_{k_1} = (p + k_1 - 1)/n_1, \quad q_{k_1} = (q + k_1 - 1)/n_1, \quad k_1 = 1, 2, \dots, n_1.$$

Since the finite product of the Mellin transform of independent random variables satisfies

$$(6) \quad \prod_{k=1}^n M_{X_k^{1/n}}(s) = M_{\prod_{k=1}^n X_k^{1/n}}(s),$$

then by (2) and (4), we have

$$(7) \quad M_X(s) = \prod_{k=1}^n M_{X_k^{1/n}}(s).$$

Thus from (6) and (7) it follows that

$$(8) \quad M_X(s) = M_{\prod_{k=1}^n X_k^{1/n}}(s).$$

If we now apply the inverse transform of the Mellin transform to both the sides of the last relation we obtain the following stochastic equality

$$(9) \quad X \stackrel{st}{=} \prod_{k=1}^n X_k^{1/n}.$$

The notation $X \stackrel{st}{=} Y$ means that X and Y have the same distribution function (d.f.).

The formula (9) means that

Theorem 1. *The second kind beta density $g(x|p, q)$ (formula (1)) of a r.v. X is equal to the density of the geometric mean of n_1 independent and nonnegative r.v.'s X_{k_1} , $k_1 = 1, \dots, n_1$, with the second kind beta distribution with the respective parameters (5).*

Now we use the Mellin transform (4) to each density of the r.v. $X_{k_1}^{1/n_1}$ for $k_1 = 1, \dots, n_1$, replacing the relevant parameters as follows

$$(10) \quad \left\{ \begin{array}{l} (p + k_1 - 1)/n_1 \text{ by } p_{k_1, k_2} = [(p + k_1 - 1)/n_1 + k_2 - 1]/n_2 \\ \text{and} \\ (q + k_1 - 1)/n_1 \text{ by } q_{k_1, k_2} = [(q + k_1 - 1)/n_1 + k_2 - 1]/n_2. \end{array} \right.$$

Then we obtain that the random variables $X_{k_1}^{1/n_1}$, $k_1 = 1, 2, \dots, n_1$, can be presented as

$$(11) \quad X_{k_1}^{1/n_1} \stackrel{st}{=} \prod_{k_2=1}^{n_2} \left(X_{k_1, k_2}^{1/n_1} \right)^{1/n_2},$$

where the random variables X_{k_1, k_2} have as their density function $\beta_2(x|p_{k_1, k_2}, q_{k_1, k_2})$.

Repeating r times the above change of parameters we get

$$(12) \quad X \stackrel{st}{=} \prod_{k_1=1}^{n_1} \prod_{k_2=1}^{n_2} \dots \prod_{k_r=1}^{n_r} X_{k_1, k_2, \dots, k_r}^{1/(n_1 n_2 \dots n_r)},$$

where the random variables X_{k_1, k_2, \dots, k_r} have densities functions $\beta_2(x|p_{k_1, k_2, \dots, k_r}, q_{k_1, k_2, \dots, k_r})$.

3. THE CASE OF INFINITE PRODUCT

Case 1.

Applying the Knar formula (Gradzstein, Ryzhik [3,8.324])

$$(13) \quad \Gamma(x+1) = 4^x \prod_{k=1}^{\infty} \left[\Gamma\left(\frac{1}{2} + \frac{x}{2^k}\right) \Gamma^{-1}\left(\frac{1}{2}\right) \right], \quad \text{Re } x > -1,$$

to the gamma functions in (2), we can rewrite (2), after some simplifications, as

$$\begin{aligned}
 M_X(s) &= \prod_{k=1}^{\infty} \left[\Gamma\left(1/2 + (p-1+s)/2^k\right) \Gamma\left(1/2 + (q-1-s)/2^k\right) \cdot \right. \\
 (14) \quad &\quad \cdot \Gamma^{-1}\left(1/2 + (p-1)/2^k\right) \Gamma^{-1}\left(1/2 + (q-1)/2^k\right) \left. \right] \\
 &= \prod_{k=1}^{\infty} h_k(s),
 \end{aligned}$$

where $-p < \text{Re } s < q$, for which we obtain that every factor $h_k(s)$ is the Mellin transform of r.v. $(X_{(k)})^{\frac{1}{2^k}}$ with the density of $X_{(k)}$ given by

$$(15) \quad \beta_2\left(x \mid \frac{1}{2} + \frac{p-1}{2^k}, \frac{1}{2} + \frac{q-1}{2^k}\right).$$

Finally, we shall prove that $\prod_1^{\infty} h_k(s)$ is the Mellin transform of $\prod_1^{\infty} (X_{(k)})^{1/2^k}$.

Proof. Since $M_X(it) = \varphi_{\ln X}(t)$, where $\varphi_X(t)$ is the characteristic function of X , we have

$$\begin{aligned}
 M_X(it) &= \varphi_{\ln X}(t) = \lim_{n \rightarrow \infty} \prod_{k=1}^n \varphi_{\ln X_{(k)}^{1/2^k}}(t) = \lim_{n \rightarrow \infty} \varphi_{\sum_1^n \ln X_{(k)}^{1/2^k}}(t) \\
 (16) \quad &= \lim_{n \rightarrow \infty} \varphi_{\ln \prod_1^n X_{(k)}^{1/2^k}}(t) = \varphi_{\ln \prod_1^{\infty} X_{(k)}^{1/2^k}}(t).
 \end{aligned}$$

Hence we get

$$(17) \quad X \stackrel{st}{=} \prod_{k=1}^{\infty} (X_{(k)})^{2^{-k}}.$$

We have proved

Theorem 2. *The second kind beta density $g(x|p, q)$ (formula (1)) is equal to the density of the infinite product of independent and non negative r.v.'s $(X_{(k)})^{2^{-k}}$ with the densities of $X_{(k)}$ determined by (15).*

Case 2. Let us now apply the generalization of Knar formula [6]

$$\begin{aligned}
 \Gamma(z+1) &= \\
 (18) \quad &= R^{Rz/(R-1)} \prod_{k=1}^{\infty} \prod_{n=1}^{R-1} \left[\Gamma\left(z/R^k + n/R\right) \Gamma^{-1}(n/R) \right]
 \end{aligned}$$

to the four gamma functions in the Mellin transform in (2); $R > 1$ is positive integer and $\text{Re}z > -1$. After some simplification we rewrite

$$\begin{aligned}
 M_X(s) &= \\
 (19) \quad &= \prod_{k=1}^{\infty} \prod_{n=1}^{R-1} \left\{ \Gamma\left[(p-1+s)/R^k + n/R\right] \Gamma\left[(q-1-s)/R^k + n/R\right] \cdot \right. \\
 &\quad \left. \cdot \Gamma^{-1}\left[(p-1)/R^k + n/R\right] \Gamma^{-1}\left[(q-1)/R^k + n/R\right] \right\} = \\
 &= \prod_{k=1}^{\infty} \prod_{n=1}^{R-1} h_{k,n}(s).
 \end{aligned}$$

Let us note that this Mellin transform differs from the product (2) by the change p, q by $p_{k,n}, q_{k,n}$ determined by

$$(20) \quad p_{k,n} = (p-1)/R^k + n/R, \quad q_{k,n} = (q-1)/R^k + n/R$$

and by coefficient at s .

Therefore every factor in right hand side of (19) is the Mellin transform of r.v. which we denote by $h_{k,n}(s)$ and we put

$$(21) \quad \prod_{n=1}^{R-1} X_{p_{k,n}, q_{k,n}}^{R-k} \stackrel{st}{=} X_{p_k, q_k}.$$

So we have proved

Theorem 3. *For each $R = 2, 3, \dots$, the density of the product of $R - 1$ r.v.'s $X_{p_{k,n}, q_{k,n}}^{R-k}$ with the densities $g(x|p_{k,n}, q_{k,n})$, where $p_{k,n}, q_{k,n}$ are determined by (20), is equal to the density of r.v. X_{p_k, q_k} .*

Finally, we obtain

$$(22) \quad X \stackrel{st}{=} \prod_{k=1}^{\infty} \prod_{n=1}^{R-1} X_{p_{k,n}, q_{k,n}}^{R-k}.$$

4. THE MODIFIED MIESHALKIN-ROGOZIN THEOREM

Our further aim is to determine an estimation of supremum of a difference between the two c.d.f. $F(x)$ and $G(x)$ if the $\sup |f(t) - g(t)|$ is known, where $f(t)$ and $g(t)$ are the characteristic functions corresponding to $F(x)$ and $G(x)$. We shall assume that $G(x)$ is the c.d.f. with the density determined by formula (1) and $F(x)$ is unknown c.d.f. which we want to estimate (see below). This problem was treated first by Gnedenko and Kolmogorov [2], where they give an estimation of such a difference using the integral $\int_{-T}^T |f(t) - g(t)|/tdt$. Next Dyson [1] showed that is not possible to determine for any $\sigma > 0$, such $\varepsilon > 0$ being dependent on σ only, that $\sup_x |F(x) - G(x)| < \sigma$ results from $\sup_x |f(t) - g(t)| < \varepsilon$. A full solution to this problem was given by Rogozin [17].

Five years later Mieshalkin and Rogozin published a paper [14] extending the results of [17]. We shall use here the Theorem 1 of the paper [14] in a somewhat modified form, concerning c.d.f.'s supported on $[0, \infty)$.

Let us assume that c.d.f. $F(x)$ and a function of bounded variation (f.b.v.) $G(x)$ fulfil the following conditions

- (C1) $F(0) = G(0) = 0$,
- (C2) $G'(x)$ exists for any $x > 0$ and $|G'(x)| < A < \infty$,
- (C3) $|f(t) - g(t)| < \varepsilon$ for $|t| < T$,

where $f(t)$ and $g(t)$ are the characteristic functions of $F(x)$ and $G(x)$, respectively.

There exists a constant C such that for $A, T, \varepsilon > 0$ and $L > 2T$ the following inequality holds

$$(23) \quad \sup_x |F(x) - G(x)| \leq C [\varepsilon \ln(LT) + A/T + \gamma(L)],$$

where

$$(24) \quad \gamma(L) = \varlimsup_{0 < x < \infty} \text{var } G(x) - \sup_x \varlimsup_{x \leq y \leq x+L} \text{var } G(y).$$

It follows from the proof of the Mieshalkin-Rogozin Theorem [14, p. 50] that

$$(25) \quad C \geq 16 \left[\ln(LT) + 2^{-3} \pi^{-1} + 1 \right]^{-1}.$$

The function $G(x)$ is c.d.f. so evidently $\text{var } G(x) = 1$ and

$$\sup_x \varlimsup_{x \leq y \leq x+L} \text{var } G(x) = \sup_x \int_x^{x+L} g(y|p, q) dy = \sup_x [G(x+L) - G(x)].$$

Let us compute the derivative

$$(26) \quad \begin{aligned} d\beta_2(x|p, q)/(dx) &= \\ &= [1/B(p, q)] (x+1)^{-2(p+q)} x^{p-2} (x+1)^{p+q-1} [-(q+1)x + p - 1]. \end{aligned}$$

It vanishes for

$$(27) \quad x_0 = (p-1)/(q+1) > 0 \text{ for } p > 1, q > 0.$$

Therefore

$$(28) \quad \sup_x \varlimsup_{x \leq y \leq x+L < \infty} G(y) = [1/B(p, q)] \sup_x \int_x^{x+L} y^{p-1} / (1+y)^{p+q} dy$$

and for the derivative of the right hand side we have also

$$(29) \quad \frac{d \varlimsup_{x \leq y \leq x+L} G(y)}{dx} = [1/B(p, q)] \left[(x+L)^{p-1} (1+x+L)^{-p-q} - x^{p-1} (1+x)^{-p-q} \right].$$

For $x < x_0$ the derivative is positive. This means that initial successive increments of the c.d.f. $G(x)$ on the segment of the lengths L increase to the size of the interval $[x_0 - L, x_0)$, where x_0 is a point in which the derivative vanishes. For $x > x_0$ the derivative is negative and vanishes when $x \rightarrow \infty$. Therefore

$$(30) \quad \sup_x \varlimsup_{x \leq y \leq x+L} G(y) = G(x_0) - G(x_0 - L)$$

and in (24) $\gamma(L)$ should be substituted by

$$(31) \quad \gamma(L) = 1 - G(x_0) + G(x_0 - L).$$

References

- [1] F.J. Dyson, *Fourier transforms of distributions functions* Canad J. Math. **5** (4) (1953), 554–558.
- [2] B.W. Gnedenko and A.N. Kolmogorov, *Rozkłady graniczne sum zmiennych losowych niezależnych*, PWN, Warszawa 1957 (in polish).
- [3] J.S. Gradstein and I.M. Ryzhik, *Tables of Integrals, Sums, Series and Products*, Moskwa 1962.
- [4] E.J. Gumbel, *The Distribution of the Range*, Biom. **36** (1962), 142.

- [5] N.L. Johnson, S. Kotz and N.B. Balakrishnan, *Continuous Univariate Distributions*, Sec. Edit. J. Wiley **1** (1995).
- [6] M. Kałuszka and W. Kryszicki, *On decompositions of some random variables*, *Metrika* **46** (2) 159–175.
- [7] M.G. Kendall and A. Stuart, *The Advanced Theory of Statistics*, Griffin Vol. 1, London.
- [8] W. Kryszicki, *et al. Probability Theory and Statistic in Problems*, Warsaw PWN, Ed V, in polish 1998.
- [9] M. Loeve, *Probability Theory*, Springer Verlag, New York, **1** (1977).
- [10] I. Lili and D. Richards, *Random discriminants*, *Ann. Statist.* **21** (1993), 1982–2001.
- [11] E. Lukacs and R.G. Laha, *Applications of characteristic functions*, Griffin 1964.
- [12] E. Lukacs, *Characteristic functions*, Griffin, Sec. Ed.
- [13] J. Marcinkiewicz and A. Zygmund, *Quelques Theoremes sur les fonctions independantes*, *Studia Mathematica* **7** (1938).
- [14] D. Mieshalkin and B.A. Rogozin, *An estimation of the distance between c.d.f.'s based on the knowledge of the absolute value of their corresponding characteristic functions and its application to the central limit theorem*, in: *The limit theorems of Probability Theory*, Uzbek. Ac. Sci., Taskent 1963 (in Russian).
- [15] A. Plucińska, *On general form of the probability density function and its application to the investigation of the distribution of rheostat resistance* *Zastosowania Matematyki* **9** (1966), 9–19.
- [16] H. Podolski, *Distribution of product and ratio of powers of independent Random Variables with the second Kind Beta Distribution*, *Scientific Bulletin of Technical University, Łódź* 1975.
- [17] B.A. Rogozin, *Some problems of the limit theorems*, p. 186–195 in: "Theory of Probability and its Application" Vol III (1958), (in Russian).
- [18] R.K. Saksena and A.M. Mathai, *Distribution of Random Variables*, *J. Roy. Statist. Soc. B.* **29** (1967), 513–525.
- [19] E. Titchmarsh, *The Theory of Functions*, 2 ed., Oxford Univ. Press 1939.

[20] B.M. Zolotarev, *The Mellin-Stieltjes Transformation in Probability Theory*, Theory Prob. Appl. **2** 433–460.

[21] Tables of integral transforms, Vol. I, Mc Graw Hill, New York 1954.

Received 10 February 2000

Revised 10 August 2000