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LIMITING PROPERTIES OF THE k -TH RECORD VALUES

1. Definitions. Let $\{X_m\}$, $m = 1, 2, \dots$, be a sequence of independent random variables with common cumulative distribution function $F(x)$, continuous with density $f(x)$. Denote by

$$(1) \quad Z_1^{(m)} \leq Z_2^{(m)} \leq \dots \leq Z_m^{(m)}$$

the order statistics in the considered sequence. First, we consider the sequence $\{Z_{m-k+1}^{(m)}\}$, $m = k, k+1, \dots$, of the k -th order statistics, where k is a constant integer. Note that the sequence $\{Z_{m-k+1}^{(m)}\}$ is non-decreasing. By elimination of repetitions in it, we get a strictly increasing subsequence which is called a *sequence of k -th record values*.

Strictly speaking, we define the indices $\{L_k(n)\}$, $n = 1, 2, \dots$, of record values in the sequence $\{Z_{m-k+1}^{(m)}\}$ by

$$L_k(1) = 1, \quad L_k(n+1) = \min\{j: Z_{L_k(n)+k-1}^{(L_k(n)+k-1)} < Z_j^{(j+k-1)}\}, \quad n = 1, 2, \dots,$$

and hence the sequence of the k -th record values by

$$(2) \quad Y_n^{(k)} = Z_{L_k(n)}^{(L_k(n)+k-1)}, \quad n = 1, 2, \dots$$

2. Distribution function of the k -th record value. Consider the sequence of random vectors

$$(3) \quad (Z_{L_k(n)}^{(L_k(n)+k-1)}, Z_{L_k(n)+1}^{(L_k(n)+k-1)}, \dots, Z_{L_k(n)+k-1}^{(L_k(n)+k-1)}), \quad n = 1, 2, \dots,$$

formed by the k last terms in sequence (1) with $m = L_k(n) + k - 1$. The first component of these vectors is the k -th record value.

LEMMA 1. *Sequence (3) is a Markov chain. The probability density $f_n(z_1, \dots, z_k)$ of vector (3) satisfies the recurrent system of equations*

$$f_1(z_1, \dots, z_k) = \begin{cases} k! f(z_1) f(z_2) \dots f(z_k), & z_1 < z_2 < \dots < z_k, \\ 0 & \text{otherwise,} \end{cases}$$

$$f_{n+1}(z_1, \dots, z_k) = \begin{cases} \sum_{i=1}^k \int_{-\infty}^{z_1} f_n(y, z_1, \dots, z_{i-1}, z_{i+1}, \dots, z_k) \frac{f(z_i)}{1-F(y)} dy, & z_1 < z_2 < \dots < z_k, \\ 0 & \text{otherwise,} \end{cases}$$

where $n = 1, 2, \dots$

THEOREM 1. *The probability density $f_n^{(k)}(y)$ of the k -th record value is of the form*

$$f_n^{(k)}(y) = \frac{k}{(n-1)!} [-k \log(1 - F(y))]^{n-1} (1 - F(y))^{k-1} f(y), \quad n = 1, 2, \dots$$

Proof. By induction, the densities $f_n(z_1, \dots, z_k)$, $n = 1, 2, \dots$, satisfy the equations

$$(4) \quad f_n(z_1, z_2, \dots, z_k) = \begin{cases} k! g_n(z_1) f(z_1) f(z_2) \dots f(z_k), & z_1 < z_2 < \dots < z_k, \\ 0 & \text{otherwise,} \end{cases}$$

where

$$g_1(z) = 1, \\ g_{n+1}(z) = k \int_{-\infty}^z g_n(y) \frac{f(y)}{1 - F(y)} dy, \quad n = 1, 2, \dots$$

It is easy to verify that

$$g_n(z) = \frac{1}{(n-1)!} [-k \log(1 - F(z))]^{n-1}, \quad n = 1, 2, \dots$$

Then, by (4), (2) and (3), we have

$$f_n^{(k)}(z) = \int_z^\infty \int_{z_2}^\infty \dots \int_{z_{k-1}}^\infty \frac{k!}{(n-1)!} [-k \log(1 - F(z))]^{n-1} f(z) f(z_2) \dots f(z_k) dz_k \dots dz_2 \\ = \frac{k}{(n-1)!} [-k \log(1 - F(z))]^{n-1} (1 - F(z))^{k-1} f(z),$$

whence the proof of the theorem is complete.

COROLLARY 1. *The probability density of the record value is of the form*

$$f_n^{(1)}(y) = \frac{1}{(n-1)!} [-\log(1 - F(y))]^{n-1} f(y)$$

(see Karlin [2]).

COROLLARY 2. *The cumulative distribution function of the k -th record value is of the form*

$$F_n^{(k)}(y) = \int_0^{-k \log(1 - F(y))} \frac{u^{n-1}}{(n-1)!} e^{-u} du, \quad n = 1, 2, \dots$$

3. Limiting distributions. While considering the limiting properties of the k -th record values we use the ideas from the papers by Tata [4] and Resnick [3], where the case $k = 1$ is considered. Here the key role

is played by the function

$$R^{(k)}(y) = -k \log(1 - F(y)).$$

THEOREM 2. *In order that there exist normalizing constants $a_n > 0$, b_n and a non-degenerate distribution function $\Phi(x)$ such that*

$$\lim_{n \rightarrow \infty} \Pr \{ Y_n^{(k)} < a_n y + b_n \} = \Phi(y) \quad \text{in continuity points of } \Phi(y)$$

it is necessary and sufficient that there exists a non-decreasing function $g(y)$ (possibly infinite valued) with more than one point of increase such that

$$\lim_{n \rightarrow \infty} (R^{(k)}(a_n y + b_n) - n) / \sqrt{n} = g(y) \quad \text{in continuity points of } g(y).$$

In addition, $\Phi(y) = N(g(y))$, where $N(x)$ is the standard normal distribution.

Proof. Since

$$\Pr \{ R^{(k)}(Y_n^{(k)}) < y \} = \int_0^y \frac{u^{n-1}}{(n-1)!} e^{-u} du,$$

we get

$$\lim_{n \rightarrow \infty} \Pr \{ R^{(k)}(Y_n^{(k)}) < \sqrt{ny} + n \} = N(y).$$

From this statement and from the equation

$$\Pr \{ Y_n^{(k)} < a_n y + b_n \} = \Pr \{ (R^{(k)}(Y_n^{(k)}) - n) / \sqrt{n} < (R^{(k)}(a_n y + b_n) - n) / \sqrt{n} \}$$

we obtain theorem 2.

Resnick [3] stated the form of the function $g(y)$ in the case $k = 1$. Analogously, for the k -th record values, it can be proved that the function $g(y)$ is of the form

$$g(y) = -\log(-\log G(y)),$$

where $G(y)$ is the cumulative distribution function belonging to one of the three types of limit distributions of extremal order statistics (see Gnedenko [1]).

THEOREM 3. *The limit distribution of the k -th record values belongs to one of the three types of distributions*

$$(5) \quad \Phi_1(y) = N(y), \quad -\infty < y < \infty,$$

$$(6) \quad \Phi_2(y) = \begin{cases} 0, & y < 0, \\ N(\log y^\alpha) & y \geq 0, \alpha > 0, \end{cases}$$

$$(7) \quad \Phi_3(y) = \begin{cases} N(\log(-y)^{-\alpha}), & y < 0, \alpha > 0, \\ 1, & y \geq 0. \end{cases}$$

Finally, we determine the domains of attraction of types (5)-(7). From the dual theorem (see [3], theorem 4.1) we obtain immediately the following

THEOREM 4. *The distribution function $F(y)$ belongs to the domain of attraction of the distribution $N(-\log(-\log G(y)))$ if and only if $1 - \exp(-\sqrt{R^{(k)}(y)})$ belongs to the domain of attraction of the limit distribution function of extremal order statistics $G(y)$.*

References

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WŁASNOŚCI GRANICZNE STATYSTYK k -REKORDOWYCH

STRESZCZENIE

W pracy podaje się definicję, rozkład prawdopodobieństwa oraz własności graniczne rozkładu prawdopodobieństwa k -tych statystyk rekordowych w ciągach niezależnych zmiennych losowych o tym samym rozkładzie.
