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A NOTE ON THE k -TH DISTANCE RANDOM VARIABLES

1. Introduction. Let $\mathbf{X}_n = (X_{n1}, X_{n2}, \dots, X_{nm})$ for $n = 0, 1, 2, \dots$ be a sequence of independent random m -dimensional vectors with the common multivariate distribution function $F(\mathbf{x}) = F(x_1, x_2, \dots, x_m)$. We write the sequence of random variables

$$Y_n = \left(\sum_{i=1}^m (X_{0i} - X_{ni})^2 \right)^{1/2} \quad (n = 1, 2, \dots).$$

Denote by $Z_1^{(n)} \leq Z_2^{(n)} \leq \dots \leq Z_n^{(n)}$ the order statistics in the sequence Y_1, Y_2, \dots, Y_n . The random variable $Z_k^{(n)}$ for $k = 1, 2, \dots, n$ will be called the k -th distance random variable.

2. Limiting distributions. Denote by $F_k^{(n)}(x) = P(Z_k^{(n)} < x)$, $x \geq 0$ ($n = 1, 2, \dots$), the distribution function of $Z_k^{(n)}$. The limiting distribution of the k -th distance random variable is defined by

$$F_k(x) = \lim_{n \rightarrow \infty} F_k^{(n)}(u_n) \quad (k = 1, 2, \dots),$$

where k is a fixed natural number and $u_n = x/n^{1/m}$, $x \geq 0$.

Kopociński in [3] and Trybuś in [5] have investigated the limiting properties of the distance random variable

$$Z_1^{(n)} = \min(Y_1, Y_2, \dots, Y_n).$$

Note that Y_n ($n = 1, 2, \dots$) is the sequence of exchangeable random variables, since

$$P(Y_{i_1} < x_{i_1}, \dots, Y_{i_r} < x_{i_r}) = \int \prod_{j=1}^r H(x_{i_j}; \mathbf{u}) dF(\mathbf{u}), \quad x_{i_j} \geq 0,$$

where

$$H(x; \mathbf{u}) = P\left(\left(\sum_{i=1}^m (u_i - X_{ni})^2 \right)^{1/2} < x \right) \quad (n = 1, 2, \dots).$$

The limiting distributions of the order statistics for exchangeable random variables have been studied among others by Berman [1] and Chanda [2].

THEOREM. *If the random vectors X_n have the bounded probability density $f(x) \leq M < \infty$, then*

$$F_k(x) = \frac{1}{(k-1)!} \int_{R^m} \left(\int_0^{F_m(x)f(u)} t^{k-1} e^{-t} dt \right) f(u) du, \quad x \geq 0,$$

where

$$V_m(x) = \frac{\pi^{m/2} x^m}{\Gamma((m/2) + 1)}$$

is the volume of an m -dimensional sphere with radius x .

Proof. It is easy to show that

$$\begin{aligned} F_k^{(n)}(x) &= \sum_{r=k}^n \sum_{1 \leq i_1 < \dots < i_r \leq n} P(Y_{i_1} < x, \dots, Y_{i_r} < x, Y_{i_{r+1}} \geq x, \dots, Y_{i_n} \geq x) \\ &= \int_{R^m} \sum_{r=k}^n \binom{n}{r} H^r(x; \mathbf{u}) (1 - H(x; \mathbf{u}))^{n-r} dF(\mathbf{u}). \end{aligned}$$

Thus

$$F_k^{(n)}(x) = \int_{R^m} F_k^{(n)}(x; \mathbf{u}) dF(\mathbf{u}),$$

where $F_k^{(n)}(x; \mathbf{u})$ is the distribution function of the k -th order statistics in the sequence of n independent random variables with the common cumulative distribution function $H(x; \mathbf{u})$.

Let η be an arbitrary positive constant and denote by $E_{\eta, n}$ the set of the points for which the oscillation of $f(x)$ in the sphere $K(x, x/n^{1/m})$ is more than η . For $x \in E_{\eta, n}^c$ we have

$$nH(u_n; x) = n \int_{|\mathbf{u}| < x/n^{1/m}} f(x + \mathbf{u}) d\mathbf{u} = V_m(x) [f(x) + \theta(x)\eta],$$

$$\text{where } |\theta(x)| \leq 1.$$

Using Theorem 4 of [4], p. 119, we obtain

$$\lim_{n \rightarrow \infty} F_k^{(n)}(u_n; x) = \frac{1}{(k-1)!} \int_0^{V_m(x)[f(x) + \theta(x)\eta]} t^{k-1} e^{-t} dt.$$

The rest of the proof is similar to the proof of Theorem 2 of [3].

COROLLARY ([3] and [5]). *The limiting distribution of the distance random variable $Z_1^{(n)}$ is of the form*

$$F_1(x) = 1 - \int_{R^m} \exp(-V_m(x)f(\mathbf{u})) f(\mathbf{u}) d\mathbf{u}.$$

References

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UWAGA O k -TYCH DYSTANSOWYCH ZMIENNYCH LOSOWYCH

STRESZCZENIE

W pracy znaleziono rozkład graniczny k -tej dystansowej zmiennej losowej, korzystając z teorii rozkładów granicznych ekstremalnych statystyk pozycyjnych. Rezultat ten jest uogólnieniem wyników Kopocińskiego [3] i Trybusia [5].
