

suffisante, mais aussi nécessaire pour l'existence du prolongement demandé, la difficulté que l'on rencontre si l'on veut (positivement) résoudre le problème P94 consiste à démontrer que la condition (C_σ) entraîne l'existence d'un module M satisfaisant à la condition $(C_\sigma-M)$.

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NOTE ON MINIMAX SOLUTIONS OF STATISTICAL
 DECISION PROBLEMS

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Given two non-void sets A and B and let $r(x, y)$ be a function defined for every $x \in A$ and $y \in B$, such that

- (1) $r(x, y) \geq 0$ for each $x \in A, y \in B$,
 (2) $\sup_{x \in A} r(x, y) < \infty$ for each $y \in B$.

It is easy to show that there exists the least σ -algebra \mathfrak{A} of subsets of A such that

- (3) $r(x, y)$ is \mathfrak{A} -measurable for each $y \in B$,
 (4) $\inf_{y \in B} r(x, y)$ is \mathfrak{A} -measurable.

Obviously, the condition (4) may be omitted if B is enumerable.

Let Ω_1 be the set of all probability measures in \mathfrak{A} . Then by (1) and (3)

$$(5) \quad 0 \leq f(\omega, y) = \int_A r(x, y) d\omega < \infty \quad \text{for each } y \in B, \omega \in \Omega_1.$$

For each pair $\omega_1, \omega_2 \in \Omega$ we shall define the function

$$(6) \quad \varrho(\omega_1, \omega_2) = \sup_{x \in \mathfrak{A}} [\omega_1(X) - \omega_2(X)].$$

It is easy to show that Ω_1 is a metric space with respect to the distance function (6). If $\Omega \subset \Omega_1$, we shall denote by $\bar{\Omega}$ the closure of Ω in Ω_1 . A probability measure which assumes only the values 0 and 1 will be called an *elementary probability measure* and the set of all elementary probability measures will be denoted by Ω_0 . If A contains at least two elements, then Ω_1 contains no isolated points, because, if $\omega \in \Omega_1$, then there exists an $\omega' \in \Omega_1$ distinct from ω , and we see at once that the sequence $\{(1 - 2^{-j})\omega(X) + 2^{-j}\omega'(X)\}_{j=1,2,3,\dots}$ of pairwise distinct points of Ω_1 converges to $\omega(X)$. Obviously, Ω_0 is isolated and each two ele-

ments of Ω_0 have the distance 1. Throughout this paper we shall always suppose that A as well as B contains at least two elements, because this assumption excludes trivial cases.

The purpose of this paper is to show that the following two theorems hold:

Theorem 1. If $\Omega_0 \subset \bar{\Omega}$ and

$$(7) \quad \sup_{x \in A} \inf_{y \in B} r(x, y) = \inf_{y \in B} \sup_{x \in A} r(x, y) = c,$$

then

$$(8) \quad \sup_{\omega \in \Omega} \inf_{y \in B} f(\omega, y) = \inf_{y \in B} \sup_{\omega \in \Omega} f(\omega, y) = c.$$

Theorem 2. If $\Omega_0 \subset \Omega$ and

$$(9) \quad \max_{x \in A} \min_{y \in B} r(x, y) = \min_{y \in B} \max_{x \in A} r(x, y) = r(x_0, y_0),$$

then there exists an $\omega^* \in \Omega$ such that

$$(10) \quad \max_{\omega \in \Omega} \min_{y \in B} f(\omega, y) = \min_{y \in B} \max_{\omega \in \Omega} f(\omega, y) = f(\omega^*, y_0) = r(x_0, y_0).$$

The application of these two theorems to statistical decision problems is based on the following interpretations of A , B , and $r(x, y)$: According to Wald's general formulation of the statistical decision problem, A is a given set of probability measures defined for all Borel sets of the finite or infinite dimensional sample space, B is a given set of decision operations, i. e. a set of transformations of the sample space on the given set of decisions, and $r(x, y)$ is the risk function, i. e. the sum of expectations of a weight function and a cost function. This formulation may be found in Wald's paper [1] and [2]. The passage from $r(x, y)$ to $f(\omega, y)$ corresponds to the randomization of the probability measures x from A . On the other hand, the randomization of the decision operations seems unreasonable from the point of view of practical applications. Roughly speaking, the theorems 1 and 2 assert that the restriction to non-randomized decision operations is legitimate under very general conditions. In particular, this simple result may be used in the decision problem in [3].

The proof of Theorem 1 will result from the following three relations:

$$(11) \quad \sup_{\omega \in \Omega} \inf_{y \in B} f(\omega, y) \leq \inf_{y \in B} \sup_{\omega \in \Omega} f(\omega, y),$$

$$(12) \quad \inf_{y \in B} \sup_{\omega \in \Omega} f(\omega, y) = \inf_{y \in B} \sup_{x \in A} r(x, y),$$

$$(13) \quad \sup_{\omega \in \Omega} \inf_{y \in B} f(\omega, y) \geq \sup_{x \in A} \inf_{y \in B} r(x, y).$$

Proof of (11). Since

$$f(\omega, y) \leq \sup_{\omega \in \Omega} f(\omega, y) \quad \text{for each } \omega \in \Omega, y \in B$$

we have

$$\inf_{y \in B} f(\omega, y) \leq \inf_{y \in B} \sup_{\omega \in \Omega} f(\omega, y) \quad \text{for each } \omega \in \Omega$$

and the last inequality implies (11).

The proof of the inequality (11) under similar conditions may be found in Wald's paper [2] and it is reproduced here only for the sake of completeness.

Proof of (12). Obviously, it is sufficient to show that

$$(14) \quad \sup_{\omega \in \Omega} f(\omega, y) = \sup_{x \in A} r(x, y) \quad \text{for each } y \in B.$$

Since

$$f(\omega, y) \leq \sup_{x \in A} r(x, y) \quad \text{for each } \omega \in \Omega, y \in B,$$

we have

$$(15) \quad \sup_{\omega \in \Omega} f(\omega, y) \leq \sup_{x \in A} r(x, y) \quad \text{for each } y \in B.$$

For each $y \in B$ and $\varepsilon > 0$ let $X_y^{(\varepsilon)}$ be the set of all $x \in A$ for which

$$r(x, y) > \sup_{x \in A} r(x, y) - \varepsilon.$$

From the properties of supremum and using (2) and (3) we have $0 \neq X_y^{(\varepsilon)} \in \mathfrak{A}$ and the integral

$$\int_{X_y^{(\varepsilon)}} r(x, y) d\omega$$

exists for each $y \in B$, $\omega \in \Omega_1$ and $\varepsilon > 0$. Furthermore, there exists an $\omega_y^{(\varepsilon)} \in \Omega_0$ such that

$$\omega_y^{(\varepsilon)}(X_y^{(\varepsilon)}) = 1 \quad \text{for each } y \in B, \varepsilon > 0,$$

because it is sufficient to choose an $x \in X_y^{(\varepsilon)}$ such that

$$\omega_y^{(\varepsilon)}(X) = \begin{cases} 1 & \text{if } x \in X, \\ 0 & \text{if } x \notin X. \end{cases}$$

Since $\Omega_0 \subset \bar{\Omega}$, therefore, to each $\delta > 0$, $\varepsilon > 0$ and $y \in B$ there exists an $\omega_{y,\delta}^{(\varepsilon)} \in \Omega$ such that

$$(16) \quad \omega_{y,\delta}^{(\varepsilon)}(X_y^{(\varepsilon)}) > 1 - \delta.$$

From the properties of supremum and using (16) we have, for each $y \in B$, $\varepsilon > 0$ and $\delta > 0$,

$$\begin{aligned} \sup_{\omega \in \bar{\Omega}} f(\omega, y) &\geq f(\omega_{y,\delta}^{(\varepsilon)}, y) = \int_A r(x, y) d\omega_{y,\delta}^{(\varepsilon)} > \\ &\geq \int_{X_y^{(\varepsilon)}} r(x, y) d\omega_{y,\delta}^{(\varepsilon)} > (1 - \delta) [\sup_{x \in A} r(x, y) - \varepsilon], \end{aligned}$$

and therefore

$$\sup_{\omega \in \bar{\Omega}} f(\omega, y) \geq \sup_{x \in A} r(x, y) \quad \text{for each } y \in B.$$

Thus, using the opposite inequality (15) we obtain (14). The relation (12) is an immediate consequence of (14).

Proof of (15). We shall first show that

$$(17) \quad \inf_{y \in B} f(\omega, y) \geq \int_A \inf r(x, y) d\omega \quad \text{for each } \omega \in \Omega.$$

Since

$$r(x, y) \geq \inf_{y \in B} r(x, y) \quad \text{for each } x \in A, y \in B,$$

we have

$$f(\omega, y) \geq \int_A \inf_{y \in B} r(x, y) d\omega \quad \text{for each } y \in B, \omega \in \Omega,$$

and (17) follows at once from this inequality. For the proof of (13) it is necessary to consider separately the following two cases:

$$(18) \quad \sup_{x \in A} \inf_{y \in B} r(x, y) = 0,$$

$$(19) \quad \sup_{x \in A} \inf_{y \in B} r(x, y) \neq 0.$$

We shall first suppose that (18) holds. Then by (17)

$$\sup_{\omega \in \bar{\Omega}} \inf_{y \in B} f(\omega, y) \geq \sup_{\omega \in \bar{\Omega}} \int_A \inf_{y \in B} r(x, y) d\omega = 0 = \sup_{x \in A} \inf_{y \in B} r(x, y),$$

i. e. (13) holds. Now let us suppose that (19) holds. Because of (2) and (11) we have

$$(20) \quad 0 < \sup_{x \in A} \inf_{y \in B} r(x, y) < \infty.$$

Before proving the inequality (15) under the condition (19) we shall first show that if the sequence $\{\omega_j(X)\}_{j=1,2,3,\dots}$ of probability measures converges to the probability measure $\omega(X)$ in the metric space Ω_1 , then

$$(21) \quad \lim_{j \rightarrow \infty} \int_A \inf_{y \in B} r(x, y) d\omega_j = \int_A \inf_{y \in B} r(x, y) d\omega.$$

Let us denote by Γ the class of all decompositions \mathbb{C} of A in a finite number of pairwise disjoint sets from \mathfrak{A} and put

$$g(\mathbb{C}, X) = \inf_{x \in X \in \mathbb{C}} \inf_{y \in B} r(x, y).$$

Then

$$\int_A \inf_{y \in B} r(x, y) d\omega = \sup_{\mathbb{C} \in \Gamma} \sum_{X \in \mathbb{C}} g(\mathbb{C}, X) \omega(X),$$

$$\int_A \inf_{y \in B} r(x, y) d\omega_j = \sup_{\mathbb{C} \in \Gamma} \sum_{X \in \mathbb{C}} g(\mathbb{C}, X) \omega_j(X) \quad \text{for } j=1,2,3,\dots,$$

and we have

$$(22) \quad \left| \int_A \inf_{y \in B} r(x, y) d\omega_j - \int_A \inf_{y \in B} r(x, y) d\omega \right| \leq \sup_{\mathbb{C} \in \Gamma} \sum_{X \in \mathbb{C}} [g(\mathbb{C}, X) |\omega_j(X) - \omega(X)|].$$

Since $\omega_j(X) - \omega(X)$ are σ -additive set functions in \mathfrak{A} , their absolute variations $\alpha_j(X)$ are measures in \mathfrak{A} for $j=1,2,3,\dots$ and it is easy to show that, for $j=1,2,3,\dots$,

$$(23) \quad |\omega_j(X) - \omega(X)| \leq \alpha_j(X) \leq 2 \sup_{Y \subset X} \sup_{Y \in \mathfrak{A}} |\omega_j(Y) - \omega(Y)|.$$

Hence, because of (23), (22) and (6)

$$(24) \quad \begin{aligned} &\left| \int_A \inf_{y \in B} r(x, y) d\omega_j - \int_A \inf_{y \in B} r(x, y) d\omega \right| \leq \\ &\leq \int_A \inf_{y \in B} r(x, y) d\alpha_j \leq \alpha_j(A) \sup_{x \in A} \inf_{y \in B} r(x, y) \leq \\ &\leq 2 \sup_{Y \in \mathfrak{A}} |\omega_j(Y) - \omega(Y)| \sup_{x \in A} \inf_{y \in B} r(x, y) = \\ &= 2\varrho(\omega_j, \omega) \sup_{x \in A} \inf_{y \in B} r(x, y). \end{aligned}$$

Since by hypothesis the sequence $\{\omega_j(X)\}_{j=1,2,3,\dots}$ converges to $\omega(X)$ in the metric space Ω_1 and (20) holds, therefore, for each $\varepsilon > 0$ there exists an index $m(\varepsilon)$ such that

$$(25) \quad \varrho(\omega_j, \omega) < \varepsilon [2 \sup_{x \in A} \inf_{y \in B} r(x, y)]^{-1} \quad \text{for each index } j > m(\varepsilon).$$

Because of (24) and (25) we have

$$\left| \int_A \inf_{y \in B} r(x, y) d\omega_j - \int_A \inf_{y \in B} r(x, y) d\omega \right| < \varepsilon \quad \text{for each index } j < m(\varepsilon),$$

i. e. (21) holds. For each $x \in A$ let ω_x be the probability measure from Ω_0 for which

$$(26) \quad \omega_x(X) = \begin{cases} 1 & \text{if } x \in X, \\ 0 & \text{if } x \text{ non-} \in X. \end{cases}$$

Since by hypothesis $\Omega_0 \subset \bar{\Omega}$, there exists a sequence

$$\{\omega_x^{(j)}(X)\}_{j=1,2,3,\dots}$$

of probability measures from Ω which converges to the probability measure $\omega_x(X)$. Obviously,

$$\sup_{\omega \in \Omega} \inf_{y \in B} f(\omega, y) \geq \inf_{y \in B} f(\omega_x^{(j)}, y) \quad \text{for each } x \in A, j=1,2,3,\dots$$

and, because of (17), (21) and (26), we have, for each $x \in A$,

$$\begin{aligned} \sup_{\omega \in \Omega} \inf_{y \in B} f(\omega, y) &\geq \liminf_{j \rightarrow \infty} \inf_{y \in B} f(\omega_x^{(j)}, y) \geq \\ &\geq \lim_{j \rightarrow \infty} \int_A \inf_{y \in B} r(x, y) d\omega_x^{(j)} = \int_A \inf_{y \in B} r(x, y) d\omega_x = \inf_{y \in B} r(x, y). \end{aligned}$$

The last relation gives at once (13) and Theorem 1 is thus proved.

The proof of Theorem 2 is very simple. By hypothesis $\Omega_0 \subset \Omega$; hence, by (26), we may substitute ω_{x_0} for ω^* . Then

$$f(\omega^*, x_0) = \int_A r(x, y_0) d\omega^* = r(x_0, y_0),$$

and Theorem 2 follows at once from Theorem 1.

Let us consider the following example: A consists of two elements 0 and 1, B is the closed interval $[0, 2]$ and $r(0, y) = y$, $r(1, y) = 2 - y$. Then \mathfrak{A} is the σ -algebra of all subsets of A . Furthermore let $\Omega = \Omega_1$. It may be seen at once that

$$\begin{aligned} \max_{x \in A} \min_{y \in B} r(x, y) &= 0 < 1 = \min_{y \in B} \max_{x \in A} r(x, y), \\ \max_{\omega \in \Omega} \min_{y \in B} f(\omega, y) &= 1 = \min_{y \in B} \max_{\omega \in \Omega} f(\omega, y). \end{aligned}$$

This example shows that the converse of Theorem 1 and 2 does not hold.

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